SOME RESULTS ON INCREMENTS OF THE PARTIALLY OBSERVED EMPIRICAL PROCESS

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ABSTRACT. The author investigates the almost sure behaviour of the increments of the partially observed, uniform empirical process. Some functional laws of the iterated logarithm are obtained for this process. As an application, new laws of the iterated logarithm are established for kernel density estimators.

1. Introduction and statement of main result

Let $\alpha_n(t) = n^{1/2}(F_n(t) - t)$ for $0 \le t \le 1$ denote the empirical uniform process, where $F_n(t) = n^{-1} \# \{U_i \le t : 1 \le i \le n\}$ is the uniform empirical distribution function, based on the first $n \ge 1$ observations from a sequence U_1, U_2, \ldots of independent uniform (0,1) random variables, defined on a probability space $(\Omega, \mathcal{A}, \mathbb{P})$. Here #A denotes the cardinality of A.

For any $0 \le a < b \le 1$, define the set $T = T_N = \{t_1, \dots, t_N\} \subseteq [0, 1]$, with $a < t_1 = t_{1,N} < \dots < t_N = t_{N,N} < b$ and $N = N_n \ge 1$.

In this paper, we investigate the almost sure behaviour of the increments of the partially observed empirical process $\{\alpha_n(t):t\in T\}$. For the statement of these results, the following notation is needed. For any $0\leq h\leq 1$ and $n\geq 1$, define the empirical increments

$$(1.1) \xi_n(h,t;s) = \alpha_n(t+sh) - \alpha_n(t), t \in [0,1-h], s \in [0,1].$$

We refer to Shorack and Wellner [21] for the basic properties of the empirical process and Deheuvels [4], Deheuvels and Révész [9], Deheuvels and Mason [5, 6, 8] among others, for a description of the behaviour of the empirical increment process.

In what follows, B[0,1] denotes the set of all bounded functions on [0,1] and AC[0,1], the set of all absolutely continuous functions $f \in B[0,1]$ of the form $f(t) = \int_0^t \dot{f}(s)ds$ for any $t \in [0,1]$ with $\int_0^1 \dot{f}^2(s)ds < \infty$. We endow B[0,1] with the topology defined by the sup-norm $||f|| = \sup_{0 \le s \le 1} |f(s)|$.

For any $c \geq 0$, denote by S_c the set of all functions of AC[0,1] satisfying $\int_0^1 \dot{f}^2(s)ds \leq c^2$ and f(0) = 0. Note that $S_c = c S_1$, where $S := S_1$ is the Strassen set (see e.g. Strassen [22]). An application of the Arzela-Ascoli theorem shows that S_c is a compact subset of $(B[0,1], \|.\|)$.

For any $A \subseteq B[0,1]$, $B \subseteq B[0,1]$, $\varepsilon > 0$, $f \in B[0,1]$ and $g \in B[0,1]$, we set $d(f,B) = \inf_{g \in B[0,1]} ||f-g||$ and $\delta(A,B) = \sup_{f \in A} d(f,B)$. The Hausdorff distance between A and B [endowed with the topology of the sup-norm], is defined

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by $\Delta(A, B) = \delta(A, B) \wedge \delta(B, A) = \inf\{\varepsilon > 0 : A \subseteq B^{\varepsilon} \text{ and } B \subseteq A^{\varepsilon}\}$. Here and elsewhere, $A^{\varepsilon} = \bigcup_{f \in A} N_{\varepsilon}(f)$ with $N_{\varepsilon}(f) = \{g \in B[0, 1] : ||f - g|| < \varepsilon\}$.

Let $\{a_n : n \geq 1\}$ be a sequence of constants satisfying

- (A1) $a_n \downarrow 0, na_n \uparrow \infty \text{ and } 0 < a_n < 1,$
- (A2) $na_n/\log n \to \infty$,
- (A3) $(\log(1/a_n))/(\log \log n) \to \infty$.

Deheuvels and Mason [6] proved the following result.

Theorem A. Under (A1)-(A3), with probability 1,

(1.2)
$$\lim_{n \to \infty} \Delta \left(\left\{ (2a_n \log(1/a_n))^{-1/2} \xi_n(a_n, t; .) : 0 \le t \le 1 - a_n \right\}, \mathcal{S} \right) = 0.$$

Our aim is to obtain the analog of Theorem A when $\{\alpha_n(t): 0 \leq t \leq 1\}$ is replaced by the partially observed empirical process. Towards this goal, a methodology exposed by Weber [26], Weber [27], Lifshits and Weber [17] is used. We refer to these articles for the study of limit laws for the partially observed Wiener process.

Let $\{h_n : n \geq 1\}$ denote a sequence of constants such that

- (H1) $h_n \downarrow 0$, $nh_n \uparrow \infty$, and $0 < h_n < 1$,
- (H2) $nh_n/\log n \to \infty$,
- (H3) $\log N_n / \log \log n \to \infty$,
- (H4) $\log N_n / \log(1/h_n) \to d \in [0, \infty].$

Remark 1.1. Here and elsewhere, we suppose the existence of a sequence of positive constants $\{\ell_n : n \geq 1\}$ such that $N_n = \lfloor \ell_n \rfloor$ for any $n \geq 1$ and $\ell_n \uparrow \infty$, $n^{-1}\ell_n \downarrow 0$ as $n \to \infty$. Here $\lfloor u \rfloor \leq u < \lfloor u \rfloor + 1$ denotes the integer part of u.

For any $n \geq 1$, set

$$(1.3) c_{n,N} = c_{n,N_n} = \{2h_n(\log N_n + \log_2 n)\}^{1/2},$$

where $\log_+ u = \log(u \vee e)$ and $\log_2 u = \log(\log_+ u)$ for $u \in \mathbb{R}$. The purpose of this paper is to establish the following theorem.

Theorem 1.1. (i) Under (H1), (H2), (H3), (H4) with $d \in [0,1)$, we have with probability 1,

(1.4)
$$\lim_{n \to \infty} \Delta(\{c_{n,N}^{-1} \, \xi_n(h_n, t; .) : t \in T_n\}, \mathcal{S}) = 0.$$

(ii) Under (H1), (H2), (H3), (H4) with $d \in [1, \infty]$, we have with probability 1,

(1.5)
$$\lim_{n \to \infty} \Delta(\{(2h_n \log(1/h_n))^{-1/2} \xi_n(h_n, t; .) : t \in T_n\}, \mathcal{S}) = 0.$$

Remark 1.2. A natural question concerns the validity of Theorem 1.1 when the technical condition (H3) is no longer satisfied. In Dindar [11], we prove that under (H1), (H2) and assumption $\log N_n/\log\log n \to c \in [0,\infty)$ as $n\to\infty$, we have with probability 1,

$$\lim_{n \to \infty} \delta(\mathcal{S}_{c/(c+1)}, \{c_{n,N}^{-1} \, \xi_n(a_n, t; .) : t \in T_n\}) = 0, \quad \text{and}$$

$$\lim_{n \to \infty} \delta(\{c_{n,N}^{-1} \, \xi_n(a_n, t; .) : t \in T_n\}, \mathcal{S}) = 0.$$

We refer to Dindar [11] for further comments and a complete proof of these last results.

2. Proofs

We postpone the proof of statements (1.4) and (1.5) until Sections 2.2 and 2.3. Below, we introduce some facts and notations suitable for our needs.

2.1. Facts and notations.

Remark 2.1. Under assumption (H3) of Theorem 1.1, it is easy to see that the sequence of norming constants $c_{n,N}$ can be replaced by $d_{n,N} = (2h_n \log N_n)^{1/2}$ for any $n \ge 1$.

Let $\{W(t): 0 \le t \le 1\}$ denote a standard Wiener process and let $\{\Pi_n(t): t \ge 0\}$ denote a sequence of right continuous Poisson processes such that $\mathbb{E}(\Pi_n(t)) = nt$ for all $n \ge 1$ and $t \ge 0$. For any $h \in (0,1)$, $s \in [0,1]$, $t \ge 0$ and integers $n \ge 1$, set

(2.1)
$$L_n(h,t;s) = n^{-1/2} \{ \Pi_n(t+sh) - \Pi_n(t) - nsh \}.$$

Fact 1. For all n sufficiently large and $1 \le m \le h_n^{-1}$, for any $t_i \in [0,1)$ such that $t_{i+1} - t_i \ge h_n$ with $i = 0, 1, \ldots, m-1$ and Borel sets $B_1, \ldots, B_m \in (B[0,1], \|.\|)$,

(2.2)
$$\mathbb{P}(\xi_n(h_n, t_i; .) \notin B_i, \ i = 0, \dots, m-1) < 2\mathbb{P}(L_n(h_n, t_i; .) \notin B_i, \ i = 0, \dots, m-1).$$

Proof. See e.g. Deheuvels and Mason [6].

Let $\{\Pi(t): t\geq 0\}$ be a standard Poisson process. We will make use of the strong approximation results of Komlós, Major and Tusnády [14, 15, 16] (see e.g. Deheuvels and Mason [7, p. 85]) to construct on the same probability space a version of the process $\{\Pi(t): t\geq 0\}$ and a standard Wiener process $\{W(t): t\geq 0\}$, such that

Fact 2. For all $T \geq 1$ and $z \in \mathbb{R}$,

(2.3)
$$\mathbb{P}(\sup_{0 \le x \le T} |\Pi(x) - x - W(x)| \ge C_1 \log T + z) \le C_2 \exp(-C_3 z),$$

for some universal constants C_1, C_2 and C_3 .

For any $f \in B[0,1]$, set

(2.4)
$$J(f) = \begin{cases} \int_0^1 (\dot{f}(s))^2 ds & \text{if } f \in AC[0,1], \\ \infty & \text{otherwise.} \end{cases}$$

For any $B \subseteq B[0,1]$, set

$$(2.5) J(B) = \inf_{f \in B} J(f).$$

Note that $J: B[0,1] \longrightarrow [0,\infty]$ is a lower semi-continuous function. For further properties, we refer to Varadhan [25], Deuschel and Stroock [10, p. 12].

For any $\lambda > 0$ and $0 \le t \le 1$, set

(2.6)
$$W_{\lambda}(t) = 2^{-1/2} \lambda^{-1} W(\lambda t).$$

We will make use of the following result stated below.

Fact 3. (i) For any closed subset F of B[0,1] endowed with the sup-norm topology,

(2.7)
$$\limsup_{\lambda \to \infty} \lambda^{-1} \log \mathbb{P}(W_{\lambda} \in F) \le -J(F).$$

(ii) For any open subset G of B[0,1] endowed with the sup-norm topology,

(2.8)
$$\limsup_{\lambda \to \infty} \lambda^{-1} \log \mathbb{P}(W_{\lambda} \in G) \ge -J(G).$$

Proof. See e.g. Schilder [20] or Deuschel and Stroock [10, Theorem 1.3.27]. \square

Introduce the Chernoff function of a standard Poisson process (see e.g. Chernoff [2] or Shorack and Wellner [21, pp. 432-433 and 856]) defined by

(2.9)
$$h(r) = \begin{cases} r \log r - r + 1 & \text{for } r > 0, \\ 1 & \text{if } r = 0, \\ \infty & \text{otherwise.} \end{cases}$$

Introduce the function

(2.10)
$$\Psi(x) = \begin{cases} 2h(1+x)/x^2 & \text{for } x \ge -1 \text{ and } x \ne 0, \\ 1 & \text{if } x = 0, \\ \infty & \text{otherwise.} \end{cases}$$

Remark 2.2. We refer to Shorack and Wellner [21, pp. 440-445] for a detailed study of Ψ properties. Among others, note that Ψ is a non-increasing function, and that $\Psi(x) \uparrow 1$ when $x \downarrow 0$.

We will need the following inequalities.

Fact 4. For any $0 < a \le 1/2$, $\lambda > 0$ and integers $n \ge 1$,

$$(2.11) \mathbb{P}\Big(\Big\|\frac{\alpha_n}{1-I}\Big\|_0^a \ge \frac{\lambda}{1-a}\Big) \le \exp\Big(-\frac{\lambda^2}{2a(1-a)}\Psi\Big(\frac{\lambda}{a\sqrt{n}}\Big)\Big).$$

Proof. See e.g. James [13] or Shorack and Wellner [21, p. 444].

Fact 5. For any $0 < a \le 1/2$, $\lambda > 0$ and integers $n \ge 1$,

(2.12)
$$\mathbb{P}(\sup_{0 \le s \le a} n |F_n(s) - s| \ge \lambda) \le 2na/\lambda^2.$$

Proof. Remark that $\{n(F_n(s) - s)/(1 - s) : 0 \le s < 1\}$ is a martingale (see e.g. Shorack and Wellner [21, p. 133]). Therefore,

$$\mathbb{P}(\sup_{0 \le s \le a} n|F_n(s) - s| \ge \lambda) \le \frac{\mathbb{E}(n^2|F_n(a) - a|^2)}{(\lambda(1 - a))^2} \le \frac{2na}{\lambda^2},$$

which concludes the proof.

2.2. Proof of Theorem 1.1 (i).

2.2.1. Preliminary lemmas. For any integers Q > 1, let

$$(2.13) d_{n,Q} := (2h_n \log_{\perp} Q)^{1/2}.$$

For any 0 < h < 1/2 and $n \ge 1$, define

(2.14)
$$\omega_{n,N}(h) = \max_{i=1,\dots,N} \sup_{0 \le s \le 1} |\alpha_n(t_i + sh) - \alpha_n(t_i)|,$$

where $t_i \in T$ for $1 \le i \le N$. We have the following lemma.

Lemma 2.1. For any 0 < h < 1/2 and $n \ge 1$,

(2.15)
$$\mathbb{P}(\omega_{n,N}(h) \ge \lambda h^{1/2}) \le N \exp\left(-\frac{\lambda^2}{2}(1-h)\Psi\left(\frac{\lambda}{\sqrt{nh}}\right)\right).$$

Proof. Statement (2.15) follows from (2.14), Fact 4 with λ replaced by $\lambda \sqrt{h}(1-h)$, a=h, and the fact that Ψ is non-increasing.

Let $\theta > 0$ whose value will be set elsewhere. For any k > 0, define

$$\mu_k = |(1+\theta)^k|.$$

We record the following result for later use.

Lemma 2.2. Under (H1), (H2), (H3), for any $\varepsilon > 0$, there exists $\theta := \theta(\varepsilon) > 0$, with $k(\varepsilon) < \infty$, such that for any $k \ge k(\varepsilon)$

$$\mathbb{P}\Big(\max_{\mu_{k-1} < n \le \mu_{k}} d_{n,N_{n}}^{-1} \omega_{n,N_{n}}(h_{n}) \ge 1 + 2\varepsilon\Big)$$

$$\le 2\mathbb{P}\Big(\frac{\omega_{\mu_{k+1},N_{\mu_{k+1}}}((1+\theta)^{2}h_{\mu_{k}})}{(1+\theta)\sqrt{h_{m}}} \ge \frac{1+\varepsilon}{1+\theta}(2\log N_{\mu_{k-1}})^{1/2}\Big).$$

Proof. This lemma is inspired by a result of Stute [23]. The proof is omitted for the sake of conciseness. We refer to Shorack and Wellner [21, pp. 545, 550-552] for a glimpse of the demonstration. \Box

We will need the following lemma.

Lemma 2.3. Under (H1), (H2), (H3), with probability 1,

(2.16)
$$\limsup_{n \to \infty} c_{n,N}^{-1} \,\omega_{n,N}(h_n) \le 1.$$

Proof. Fix $\varepsilon > 0$ and consider the event

$$E_n = \left\{ \frac{\omega_{n,N_n}(h_n)}{\sqrt{h_n}} \ge (1+\varepsilon)(2\log N_n)^{1/2} \right\}.$$

Assertion (2.16) is verified as soon as $\sum_n \mathbb{P}(E_n) < \infty$. Hence using Lemma 2.2 and the Borel-Cantelli lemma, it suffices to show that $\sum_k \mathbb{P}(F_k) < \infty$, where

$$F_k = \left\{ \frac{\omega_{\mu_{k+1}, N_{\mu_{k+1}}} ((1+\theta)^2 h_{\mu_k})}{(1+\theta)\sqrt{h_{\mu_k}}} \ge \frac{1+\varepsilon}{1+\theta} (2\log N_{\mu_{k-1}})^{1/2} \right\}.$$

For any $m \geq 1$, let $\lambda_m = (2 \log N_m)^{1/2}$. We use Lemma 2.1 with λ replaced by $(1+\varepsilon)(1+\theta)^{-1}\lambda_{\mu_{k-1}}$ and $h = (1+\theta)\sqrt{h_{\mu_k}}$ to obtain

$$\mathbb{P}(F_k) \le N_{\mu_{k+1}} \exp\Big\{ -\frac{1}{2} \Big(\frac{1+\varepsilon}{1+\theta} \Big)^2 \lambda_{\mu_{k-1}}^2 \Big(1 - (1+\theta)^2 h_{\mu_k} \Big) \\ \times \Psi\Big(\frac{1+\varepsilon}{(1+\theta)^2} \frac{\lambda_{\mu_{k-1}}}{\sqrt{\mu_{k+1} h_{\mu_k}}} \Big) \Big\}.$$
(2.17)

Consider $\theta > 0$ such that $\{(1+\varepsilon)/(1+\theta)\}^2(1-(1+\theta)^2h_{\mu_k}) > 1+\varepsilon/4$. This, combined with (2.17) entails

$$\mathbb{P}\left(\frac{\omega_{\mu_{k+1},N_{\mu_{k+1}}}((1+\theta)^{2}h_{\mu_{k}})}{\sqrt{h_{\mu_{k}}}} \ge (1+\varepsilon)\lambda_{\mu_{k-1}}\right) \\
\le N_{\mu_{k+1}} \exp\left\{-\frac{1}{2}(1+\frac{\varepsilon}{4})\lambda_{\mu_{k-1}}^{2}\Psi\left(\frac{1+\varepsilon}{(1+\theta)^{2}}\frac{\lambda_{\mu_{k-1}}}{\sqrt{\mu_{k+1}h_{\mu_{k}}}}\right)\right\}.$$

By (H1)-(H2) and Remark 1.2, we obtain for any k sufficiently large

(2.19)
$$\Psi\left(\frac{1+\varepsilon}{(1+\theta)^2} \frac{\lambda_{\mu_{k-1}}}{\sqrt{\mu_{k+1} h_{\mu_k}}}\right) > 1-\varepsilon',$$

where $\varepsilon' > 0$ is such that $(1 + \varepsilon/4)(1 - \varepsilon') > 1 + \varepsilon/16$. This, combined with (2.19) implies

(2.20)
$$\mathbb{P}\left(\frac{\omega_{\mu_{k+1},N_{\mu_{k+1}}}((1+\theta)^{2}h_{\mu_{k}})}{\sqrt{h_{\mu_{k}}}} \ge (1+\varepsilon)\lambda_{\mu_{k-1}}\right) \le N_{\mu_{k+1}} \exp\left\{-(1+\frac{\varepsilon}{16})\log N_{\mu_{k-1}}\right\}.$$

The definition of μ_k and a few computation prove that the RHS of (2.20) is bounded by k^{-2} for any k sufficiently large. Therefore, $\sum_k \mathbb{P}(F_k) < \infty$ and (2.16) is proved.

Let $\gamma > 0$ be a constant whose value will be set elsewhere. For any $k \geq 1$, set

$$(2.21) n_k = \lfloor (1+\gamma)^k \rfloor.$$

We have the following result.

Lemma 2.4. For any k sufficiently large,

- (1i) $\log n_k \sim k \log(1+\gamma)$,
- (1ii) $\log_2 n_k \sim \log k$,
- (1iii) $\log_2 n_k \sim \log_2 n_{k+1}$,
- (1iv) $n_{k+1} n_k = (1 + o(1))\gamma n_k$

Proof. Proofs of (1i)-(1iv) are trivial and therefore omitted.

For any $\varepsilon > 0$ and $k \ge 1$, consider the sets

$$C_k(\varepsilon, \gamma) = \{ (n/n_{k+1})^{1/2} d_{n_{k+1}, N_{n_{k+1}}}^{-1} \xi_n(h_{n_{k+1}}, t_i; .) \notin \mathcal{S}^{\varepsilon},$$
for $i \in \{1, ..., N_{n_{k+1}}\}$ and $n_k < n \le n_{k+1} \}$

and
$$D_k(\varepsilon, \gamma) = \{d_{n_{k+1}, N_{n_{k+1}}}^{-1} \xi_{n_{k+1}}(h_{n_{k+1}}, t_i; .) \notin \mathcal{S}^{\varepsilon}, \text{ for } i \in \{1, \dots, N_{n_{k+1}}\} \}.$$

Lemma 2.5. For any $\varepsilon > 0$ and $\gamma > 0$, there exists a $k(\varepsilon, \gamma) < \infty$ such that, for any $k \geq k(\varepsilon, \gamma)$

(2.22)
$$\mathbb{P}(C_k(\varepsilon,\gamma)) \le 2\mathbb{P}(D_k(\varepsilon/2,\gamma)).$$

Proof. See e.g. Deheuvels and Mason [6]. Define the conditions

$$E_{k,i,n}(\varepsilon) = \left\{ \left(\frac{n}{n_{k+1}} \right)^{1/2} d_{n_{k+1},N_{n_{k+1}}}^{-1} \xi_n(h_{n_{k+1}}, t_i; .) \notin \mathcal{S}^{\varepsilon} \right\},$$

$$E_{k,n} = \bigcup_{i=1}^{N_{n_{k+1}}} E_{k,i,n}(\varepsilon),$$

$$F_{k,i,n}(\varepsilon) = \big\{ d_{n_{k+1},N_{n_{k+1}}}^{-1} \, \big\| \xi_{n_{k+1}}(h_{n_{k+1}},t_i;.) - \big(\frac{n}{n_{k+1}}\big)^{1/2} \xi_n(h_{n_{k+1}},t_i;.) \big\| < \varepsilon \big\}.$$

For any $\varepsilon' > 0$, $\varepsilon'' > 0$, the conditions $\{E_{k,i,n}(\varepsilon') : i \geq 1\}$ and $\{F_{k,i,n}(\varepsilon'') : i \geq 1\}$ are independent. Therefore,

$$\mathbb{P}(C_k(\varepsilon,\gamma)) = \sum_{q=n_k+1}^{n_{k+1}} \sum_{i=1}^{N_{n_{k+1}}} \mathbb{P}(E_{k,i,q}(\varepsilon) \cap_{j=1}^{i-1} \overline{E}_{k,j,q}(\varepsilon) \cap_{r=n_k+1}^{q-1} \overline{E}_{r,q}(\varepsilon)).$$

Here and elsewhere, \overline{A} denotes the closure of the set A. Moreover, observe that

$$\inf_{n_{k} < n \le n_{k+1}} \inf_{1 \le l \le N_{n_{k+1}}} \mathbb{P}(F_{k,l,n}(\varepsilon/2)) \mathbb{P}(C_{k}(\varepsilon,\gamma))$$

$$\le \sum_{q=n_{k}+1}^{n_{k+1}} \sum_{i=1}^{N_{n_{k+1}}} \mathbb{P}(E_{k,i,q}(\varepsilon) \cap F_{k,i,q}(\varepsilon/2) \cap_{j=1}^{i-1} \overline{E}_{k,j,q}(\varepsilon) \cap_{r=n_{k}+1}^{q-1} \overline{E}_{r,q}(\varepsilon))$$

$$(2.23) \qquad \le \mathbb{P}(\bigcup_{i=1}^{N_{n_{k+1}}} E_{k,i,n_{k+1}}(\varepsilon/2)) = \mathbb{P}(D_{k}(\varepsilon/2,\gamma)).$$

On the other hand, we have

$$\mathbb{P}(\overline{F}_{k,l,n}(\varepsilon/2)) = \mathbb{P}(n_{k+1}^{-1/2} \sup_{0 \le u \le h_{n_{k+1}}} (n_{k+1} - n) | F_{n_{k+1} - n}(u) - u | \ge \frac{\varepsilon}{2} d_{n_{k+1}, N_{n_{k+1}}}).$$

Next, use Fact 5 and (2.12) to obtain

$$(2.24) \mathbb{P}(\overline{F}_{k,l,n}(\varepsilon/2)) \le 8 \frac{(n_{k+1} - n_k)h_{n_{k+1}}}{\varepsilon^2 n_{k+1} d_{n_{k+1},N_{n_{k+1}}}^2}.$$

Statement (1.3) and (H3) jointly imply that the RHS of (2.24) converges to 0. Hence, $(\inf_{n_k < n \le n_{k+1}} \inf_{1 \le l \le N_{n_{k+1}}} \mathbb{P}(F_{k,l,n}(\varepsilon/2)))^{-1} \to 1$ as $k \to \infty$. This, combined with (2.23), entails (2.22).

The blocking arguments described in the next two lemmas will be suitable for the proof of (1.4).

Lemma 2.6. Under (H1), (H2), (H3), with probability 1,

(2.25)
$$\lim_{\gamma \downarrow 0} \limsup_{k \to \infty} \left\{ \max_{n_k < n \le n_{k+1}} \max_{1 \le i \le N_{n_{k+1}}} \left| \left(\frac{n}{n_{k+1}} \right)^{1/2} d_{n_{k+1}, N_{n_{k+1}}}^{-1} - d_{n, N_n}^{-1} \right| \times \left\| \xi_n(h_{n_{k+1}}, t_i; .) \right\| \right\} = 0.$$

Proof. Observe that

$$(2.26) \qquad \left| \left(\frac{n}{n_{k+1}} \right)^{1/2} d_{n_{k+1}, N_{n_{k+1}}}^{-1} - d_{n, N_n}^{-1} \right| = d_{n, N_n}^{-1} \left| \left(\frac{n}{n_{k+1}} \right)^{1/2} \frac{d_{n, N_n}}{d_{n_{k+1}, N_{n+1}}} - 1 \right|.$$

By (H1), it is easy to see that $\sqrt{n}d_n \uparrow \infty$ as $n \to \infty$. Therefore, for any k sufficiently large,

$$\left| \left(\frac{n}{n_{k+1}} \right)^{1/2} \frac{d_{n,N_n}}{d_{n_{k+1},N_{n_{k+1}}}} - 1 \right| \le \left| \left(\frac{n_k}{n_{k+1}} \right)^{1/2} \frac{d_{n_k,N_{n_k}}}{d_{n_{k+1},N_{n_{k+1}}}} - 1 \right|.$$

Observe that (H1) and (2.21) imply for all k large enough

(2.28)
$$\frac{1}{1+\gamma} \le \frac{n_k h_{n_k}}{n_{k+1} h_{n_{k+1}}} \le 1.$$

Moreover, (H1), (H3) and (2.21) combined with Remark 1.1 entails for any k large enough

(2.29)
$$1 - \frac{\log(1+\gamma)}{4\log k} \le \frac{\log N_{n_k}}{\log N_{n_{k+1}}} \le 1.$$

Therefore, (2.28) and (2.29) jointly imply that the RHS of (2.26) converges to 0 as $k \to \infty$ and $\gamma \downarrow 0$. Next, we use (H1) and (2.14) to obtain for any k large

enough and $1 \leq i \leq N_{n_{k+1}}$, that $\|\xi_n(h_{n_{k+1}}, t_i; .)\| \leq \omega_{n, N_{n_{k+1}}}(h_n)$. Moreover, for any $n_k < n \leq n_{k+1}$, we have $d_{n, N_{n_{k+1}}} d_{n, N_n}^{-1} \to 1$ as $k \to \infty$. These last statements, combined with Lemma 2.3 entail (2.25).

Lemma 2.7. Under (H1), (H2), (H3), we have with probability 1,

(2.30)

$$\lim_{\gamma \downarrow 0} \limsup_{k \to \infty} \left\{ \max_{n_k < n \le n_{k+1}} \max_{1 \le i \le N_{n_{k+1}}} d_{n,N_n}^{-1} \| \xi_n(h_{n_{k+1}}, t_i; .) - \xi_n(h_n, t_i; .) \| \right\} = 0.$$

Proof. For any $n_k < n \le n_{k+1}$ and $1 \le i \le N_{n_{k+1}}$, we have

Moreover, for any $n_k < n \le n_{k+1}$ and any k large enough, (H1) implies

$$(2.32) h_n - h_{n_{k+1}} \le h_n \left(1 - \frac{h_{n_{k+1}}}{h_{n_k}}\right) \le \gamma h_n.$$

From Lemma 2.3, we get with probability 1

(2.33)
$$\limsup_{k \to \infty} d_{n, N_{n_{k+1}}}^{-1} \omega_{n, N_{n_{k+1}}}(\gamma h_n) \le \gamma^{1/2}.$$

Next, observe that for any $n_k < n \le n_{k+1}$, $d_{n,N_{n_{k+1}}} d_{n,N_n}^{-1} \longrightarrow 1$ as $k \to \infty$. This last statement, combined with (2.31)-(2.33) entails (2.30) as $k \downarrow 0$.

Proof of Theorem 1.1 (i) is divided in two steps.

2.2.2. Proof of Theorem 1.1 (i). First step. As a first step, we show that, for any $\varepsilon > 0$, there exists with probability 1 an $n(\varepsilon) < \infty$ such that, for any $n \ge n(\varepsilon)$

(2.34)
$$\mathcal{S} \subseteq \{d_{n,N}^{-1} \, \xi_n(h_n, t; .) : t \in T_n\}^{\varepsilon}.$$

Recall that S is a compact subset of $(B[0,1], \|.\|)$ (cf. Conway [3, p. 177]). Hence, for any $\varepsilon > 0$, we can choose an integer $m \ge 1$ and $f_i \in S$, $1 \le i \le m$, such that $S \subseteq \bigcup_{i=1}^m N_{\varepsilon}(f_i)$. Therefore, (2.34) will be verified as soon as we show that for any $f \in S$ and $\varepsilon > 0$, we have

(2.35)
$$\sum_{n} P_n(f,\varepsilon) := \sum_{n} \mathbb{P}(\bigcap_{i=1}^{N_n} d_{n,N_n}^{-1} \, \xi_n(h_n,t_i;.) \notin N_{\varepsilon}(f)) < \infty,$$

where $t_i \in T$ for $i = 1, \ldots, N_n$.

By Fact 1, it is easy to show that

$$(2.36) P_n(f,\varepsilon) \le 2 \left\{ 1 - \mathbb{P} \left((2nh_n \log N_n)^{-1/2} (\Pi(nh_n) - nh_n) \in N_{\varepsilon}(f) \right) \right\}^{N_n}$$

Set $P_n^{(1)}(f,\varepsilon) := \mathbb{P}(\{2nh_n \log N_n\}^{-1/2}\{\Pi(nh_n.) - nh_n.\} \in N_{\varepsilon}(f))$ and let $\{W(t) : t \geq 0\}$ denote a standard Wiener process such that (2.3) of Fact 2 is verified. For any $n \geq 1$, we have

$$(2.37) P_n^{(1)}(f,\varepsilon) \ge \mathbb{P}\left(\frac{W(nh_n.)}{(2nh_n\log N_n)^{1/2}} \in N_{\varepsilon/2}(f)\right) - \mathbb{P}\left(\sup_{0 \le s \le 1} \frac{|\Pi(nh_ns) - nh_ns - W(nh_ns)|}{(2nh_n\log N_n)^{1/2}} \ge \frac{\varepsilon}{2}\right).$$

Set

(2.38)
$$P_n^{(2)}(f,\varepsilon) := \mathbb{P}\left(\frac{W(nh_n)}{(2nh_n \log N_n)^{1/2}} \in N_{\varepsilon/2}(f)\right)$$
$$= \mathbb{P}\left(\frac{W(\log N_n)}{2^{1/2} \log N_n} \in N_{\varepsilon/2}(f)\right).$$

Note that $N_{\varepsilon/2}(f)$ is an open set of B[0,1] [endowed with the sup-norm topology]. Therefore, Fact 3 and (2.8) imply for any n sufficiently large

$$(2.39) P_n^{(2)}(f,\varepsilon) \ge \exp(-(1-\delta)\log N_n),$$

where $\delta \in (0,1)$ verifies $J(N_{\varepsilon/2}(f)) < 1 - \delta < 1$. Now, let us define

(2.40)
$$P_n^{(3)} := \mathbb{P}\left(\sup_{0 < s < 1} \frac{|\Pi(nh_n s) - nh_n s - W(nh_n s)|}{(2nh_n \log N_n)^{1/2}} \ge \frac{\varepsilon}{2}\right).$$

By Fact 2 and (H1)-(H3), we get for any n large enough

$$P_n^{(3)} \le \mathbb{P}\left(\sup_{0 \le s \le nh_n} |\Pi(s) - s - W(s)| \ge \frac{\varepsilon}{4} (2nh_n \log N_n)^{1/2} + C_1 \log(nh_n)\right)$$

$$(2.41) \leq C_2 \exp\left(-C_3 \frac{\varepsilon}{4} (2nh_n \log N_n)^{1/2}\right) \leq \frac{1}{2} \exp\left(-(1-\delta) \log N_n\right).$$

We jointly use (2.37), (2.39) and (2.41) to obtain

$$(2.42) (1 - P_n^{(1)}(f, \varepsilon))^{N_n} \le \exp\left(-\frac{1}{2} \frac{N_n^{\delta}}{(\log n)^{1-\delta}}\right).$$

By (H3), for any integer r > 1 and any n large enough, we have

$$(2.43) -\frac{N_n^{\delta}}{(\log n)^{1-\delta}} < -(\log n)^r.$$

Therefore, for any n sufficiently large

$$(2.44) (1 - P_n^{(1)}(f, \varepsilon))^{N_n} \le \exp\left(-\frac{1}{2}(\log n)^r\right).$$

This last statement implies $\sum_{n} P_n(f, \varepsilon) < \infty$. Hence, the Borel-Cantelli lemma entails (2.34).

2.2.3. Proof of Theorem 1.1 (i). Second step. As a second step we prove that, with probability 1

(2.45)
$$\lim_{n \to \infty} \delta(\{d_{n,N}^{-1} \xi_n(h_n, t; .) : t \in T_n\}, \mathcal{S}) = 0.$$

The successive use of Lemma 2.5, the Borel-Cantelli lemma, and Lemmas 2.6 and 2.7 show that (2.45) is true as soon as we have, for any $\varepsilon > 0$ and $\gamma > 0$

(2.46)
$$\sum_{k} \mathbb{P}(D_k(\varepsilon, \gamma)) < \infty.$$

For any $k \geq 1$, observe that

$$\mathbb{P}(d_{n_{k+1},N_{n_{k+1}}}^{-1}\xi_{n_{k+1}}(h_{n_{k+1}},t;.) \notin \mathcal{S}^{\varepsilon} \text{ for a } t \in T)$$

$$(2.47) \leq N_{n_{k+1}} \mathbb{P}(d_{n_{k+1}, N_{n_{k+1}}}^{-1} \xi_{n_{k+1}}(h_{n_{k+1}}, 0; .) \notin \mathcal{S}^{\varepsilon}).$$

Set

(2.48)
$$Q_k^{(0)} := \mathbb{P}(d_{n_{k+1}, N_{n_{k+1}}}^{-1} \xi_{n_{k+1}}(h_{n_{k+1}}, 0; .) \notin \mathcal{S}^{\varepsilon}).$$

We use Fact 1 to obtain for any $k \ge 1$,

$$(2.49) \quad Q_k^{(0)} \leq 2\mathbb{P}(d_{n_{k+1},N_{n_{k+1}}}^{-1}L_{n_{k+1}}(h_{n_{k+1}},0;.) \notin \mathcal{S}^{\varepsilon})$$

$$\leq 2\mathbb{P}\Big(\frac{W(n_{k+1}h_{n_{k+1}})}{(2n_{k+1}h_{n_{k+1}}\log N_{n_{k+1}})^{1/2}} \notin \mathcal{S}^{\varepsilon/2}\Big)$$

$$+ 2\mathbb{P}\Big(\sup_{0\leq s\leq 1} \frac{|\Pi(n_{k+1}h_{n_{k+1}}s) - n_{k+1}h_{n_{k+1}}s - W(n_{k+1}h_{n_{k+1}}s)|}{(2n_{k+1}h_{n_{k+1}}\log N_{n_{k+1}})^{1/2}} \geq \frac{\varepsilon}{2}\Big).$$

Now consider

$$(2.50) Q_k^{(1)} := \mathbb{P}\Big(\frac{W(n_{k+1}h_{n_{k+1}})}{(2n_{k+1}h_{n_{k+1}}\log N_{n_{k+1}})^{1/2}} \in B[0,1] \setminus \mathcal{S}^{\varepsilon/2}\Big).$$

Observe that $B[0,1] \setminus S^{\varepsilon/2}$ is a closed subset of B[0,1] [endowed with the sup-norm topology]. Therefore, Fact 3 and (2.7) entail for any k large enough

(2.51)
$$Q_k^{(1)} \le \exp(-(1+\delta)\log N_{n_{k+1}}),$$

where $\delta > 0$ is such that $1 < 1 + \delta < J(B[0,1] \setminus \mathcal{S}^{\varepsilon/2})$. Let

$$(2.52) \quad Q_k^{(2)} := \mathbb{P}\Big(\sup_{0 \le s \le 1} \frac{|\Pi(n_{k+1}h_{n_{k+1}}s) - n_{k+1}h_{n_{k+1}}s - W(n_{k+1}h_{n_{k+1}}s)|}{(2n_{k+1}h_{n_{k+1}}\log N_{n_{k+1}})^{1/2}} \ge \frac{\varepsilon}{2}\Big).$$

By (H1)-(H2) and Fact 2, we have for any k large enough

$$Q_k^{(2)} \le C_2 \exp\left(-C_3 \frac{\varepsilon}{2} (2n_{k+1} h_{n_{k+1}} \log N_{n_{k+1}})^{1/2}\right)$$

$$(2.53) \qquad \le \exp(-(1+\delta) \log N_{n_{k+1}}).$$

Then (H3) and statements (2.21), (2.48), (2.49), (2.51), (2.53) jointly imply

$$(2.54) N_{n_{k+1}}Q_k^{(0)} \le 4 \frac{1}{N_{n_{k+1}}^{\delta}(\log n_{k+1})^{1+\delta}}$$

$$\le \frac{2^{1+2\delta}}{(\log(1+\gamma))^{1+2\delta}} \frac{1}{(k+1)^{1+2\delta}}.$$

Hence, we have $\sum_{k} N_{n_{k+1}} Q_k^{(0)} < \infty$. Now, use Lemma 2.5, the Borel-Cantelli lemma and Lemmas 2.6 and 2.7 to obtain (2.45).

By (2.34) and (2.45) we conclude that (1.4) is verified. This achieves the proof of Theorem 1.1 (i).

2.3. **Proof of Theorem 1.1 (ii).** The proof of (1.5) is derived from Theorem A of Deheuvels and Mason [6] since (1.2) holds for any subset $T \subseteq [0, 1 - h_n]$, finite or infinite. Therefore a detailed proof of (1.5) is omitted.

Remark 2.3. Using the techniques of paragraph 2.2, with a particular set T_n of equidistant points, we prove that the sequence of norming constants $c_{n,N}$ is optimal.

3. Applications

In this section, we establish new laws of iterated logarithm for kernel density estimators.

Let $\{X_n : n \geq 1\}$ denote a sequence of independent and identically distributed random variables defined on $(\Omega, \mathcal{A}, \mathbb{P})$ with distribution function G. We suppose that G is a continuous function on [A, B], $-\infty \leq A < B \leq \infty$, with derivative

g(x) = G'(x), positive and continuous on [A, B]. Let C and D denote two constants such that A < C < D < B.

Assume $N \geq 1$ and consider the points x_1, \ldots, x_N of [A, B] such that

$$A < C < x_1 = x_{1,N} = x_{1,N_n} < \ldots < x_N = x_{N,N} = x_{N_n,N_n} < D < B.$$

Set $R = R_N = \{x_1, \dots, x_N\}$ and let K be a function satisfying the following assumptions,

- (K1) K is of bounded variation on $(-\infty, \infty)$,
- (K2) For $M \in (0, \infty)$, K(u) = 0 for any $|u| \ge M$,
- (K3) $\int_{-\infty}^{\infty} K(u)du = 1.$

We introduce the kernel density estimator of Parzen [18] and Rosenblatt [19] (see also Bosq and Lecoutre [1]), defined for any $x \in [A, B]$ by

(3.1)
$$f_n(x) = (nh_n)^{-1} \sum_{i=1}^n K\left(\frac{x - X_i}{h_n}\right),$$

where $\{h_n : n \ge 1\}$ and $\{N_n : n \ge 1\}$ follow the assumptions (H1)-(H4).

Remark 3.1. Properties of f_n can be found in Bosq and Lecoutre [1]. See also Stute [24], Deheuvels and Mason [6] among others, for the study of limit laws of the iterated logarithm for the kernel density estimator.

Theorem 3.1. (i) Under (H1), (H2), (H3), (H4) with $d \in [0,1)$ and (K1)-(K3), with probability 1,

(3.2)
$$\limsup_{n \to \infty} \left(\frac{nh_n}{2 \log N_n} \right)^{1/2} \max_{x \in \{x_1, \dots, x_{N_n}\}} \pm \left(\frac{f_n(x) - \mathbb{E}f_n(x)}{\sqrt{g(x)}} \right)$$
$$= \left(\int_{-\infty}^{\infty} K^2(u) du \right)^{1/2}.$$

(ii) Under (H1), (H2), (H3), (H4) with $d \in [1, \infty]$ and (K1)-(K3), with probability 1.

(3.3)
$$\lim_{n \to \infty} \left(\frac{nh_n}{2\log(1/h_n)} \right)^{1/2} \max_{x \in \{x_1, \dots, x_{N_n}\}} \pm \left(\frac{f_n(x) - \mathbb{E}f_n(x)}{\sqrt{g(x)}} \right)$$
$$= \left(\int_{-\infty}^{\infty} K^2(u) du \right)^{1/2}.$$

Remark 3.2. Theorem 3.1 may be purposeful for experimental situations where we recall that $f_n(x)$ is unknown continuously over [A, B], but only on a discrete subset of [A, B]. We can also justify the use of Theorem 3.1 when prebinning methods are considered (see e.g. Härdle [12]).

Proof of Theorem 3.1. Assume that M=1/2 and let $\widetilde{K}(u):=K(1/2-u)$ for $u\in[0,1]$. Next, consider

(3.4)
$$\widetilde{f}_n(x) = (nh_n)^{-1} \sum_{i=1}^n \widetilde{K}\left(\frac{X_i - x}{h_n}\right),$$

and observe that

$$\widetilde{f}_n(x) = f_n(x + \frac{1}{2}h_n).$$

Proof of (i). The use of (H1) and (K2), combined with the continuity of g on [A, B] implies that (3.2) is true whenever we prove with probability 1,

(3.6)
$$\limsup_{n \to \infty} \left(\frac{nh_n}{2\log N_n} \right)^{1/2} \max_{x \in \{x_1, \dots, x_{N_n}\}} \pm \left(\frac{\widetilde{f}_n(x) - \mathbb{E}\widetilde{f}_n(x)}{\sqrt{g(x)}} \right) = \left(\int_0^1 \widetilde{K}^2(u) du \right)^{1/2}.$$

For any $x \in R$, (3.4) entails

$$\widetilde{f}_n(x) = \int_0^1 h_n^{-1} \widetilde{K}(u) dG_n(x + h_n u)$$

$$= -\int_0^1 h_n^{-1} (G_n(x + h_n u) - G_n(x)) d\widetilde{K}(u),$$

where G_n is the empirical distribution function based on X_1, \ldots, X_n . Next using the equality $G_n(x) = F_n(G(x))$ (see e.g. Shorack and Wellner [21, pp. 3-5]), we see that

$$(3.7) n^{1/2}h_n(\widetilde{f}_n(x) - \mathbb{E}\widetilde{f}_n(x)) = -\int_0^1 \left(\alpha_n(G(x+h_nu)) - \alpha_n(G(x))\right)d\widetilde{K}(u).$$

Let $\delta_0 > 0$ be a constant whose definition will be given elsewhere. Consider the subset $\widetilde{R} = \widetilde{R}_N = \widetilde{R}_{N_n} = \{\widetilde{x}_1, \dots, \widetilde{x}_{k_{N_n}}\} \subseteq R$, such that for any n sufficiently large we have $\widetilde{x}_{k_{N_n}} - \widetilde{x}_1 < \delta_0$.

Next, we define the dual set of \widetilde{R} , $\widetilde{T} = \widetilde{T}_N = \widetilde{T}_{N_n} = \{\widetilde{t}_1, \dots, \widetilde{t}_{k_{N_n}}\} \subseteq T'$, with $G(\widetilde{x}_1) = \widetilde{t}_1 < \dots < G(\widetilde{x}_{k_{N_n}}) = \widetilde{t}_{k_{N_n}}$. For any $n \geq 1$, define the sequence of constants $a_n = \Lambda h_n$, where $\Lambda > 0$ will be described elsewhere. It is easy to see that

(3.8)
$$\{ (2a_n \log(N_n/\Lambda))^{-1/2} \xi_n(a_n, t; .) : t \in \widetilde{T} \}$$

$$= \{ (2a_n \log(N_n/\Lambda))^{-1/2} (\alpha_n(G(x) + \Lambda h_n.) - \alpha_n(G(x))) : x \in \widetilde{R} \}.$$

Moreover $\{a_n : n \ge 1\}$ verifies assumptions (H1), (H2), (H3), (H4) with $d \in [0, 1)$. Hence, for $\varepsilon > 0$, statement (1.4) implies with probability 1 for any n sufficiently large,

(3.9)
$$\left\{ (2a_n \log(N_n/\Lambda))^{-1/2} \xi_n(a_n, t; .) : t \in \widetilde{T} \right\} \subseteq \mathcal{S}^{\varepsilon/8}.$$

Next, we set $\Lambda = \max_{1 \leq i \leq k_N} g(x_i)$. For any $x \in \widetilde{R}$, $u \in [0, 1]$, define

(3.10)
$$\Theta_n(x;u) := \frac{G(x + h_n u) - G(x)}{\Lambda h_n}.$$

Recall that g is continuous on [A, B]. Hence, we choose $\delta_0 > 0$ such that for any n large enough and $x \in \widetilde{R}$,

$$|\Lambda^{-1}g(x) - 1| < \frac{\varepsilon^2}{32}.$$

Moreover, the continuity of G, (3.10) and (3.11) jointly imply for any n large enough, $u \in [0, 1]$ and $x \in \widetilde{R}$,

$$(3.12) |\Theta_n(x;u) - u| \le \frac{\varepsilon^2}{16}.$$

For any $t \in \widetilde{T}$ and $u \in [0,1]$, define $\Xi(u) := (2a_n \log(N_n/\Lambda))^{-1/2} \xi_n(a_n,t;u)$ and observe that for any $f \in B[0,1]$,

$$(3.13) \qquad (2a_n \log(N_n/\Lambda))^{-1/2} \left(\alpha_n(G(x+h_n.)) - \alpha_n(G(x))\right) - f \|$$

$$\leq \|\Xi(\Theta_n(x;.)) - \Xi(.)\|$$

$$+ \|(2a_n \log(N_n/\Lambda))^{-1/2} \left(\alpha_n(G(x) + \Lambda h_n.) - \alpha_n(G(x))\right) - f \|.$$

By (3.9), observe that $\Xi \in \mathcal{S}^{\varepsilon/8}$. Therefore, a simple computation prove that $\|\Xi(\Theta_n(x;.)) - \Xi(.)\| \le \varepsilon/2$. Moreover for any $f \in \mathcal{S}$, (3.9) implies

$$\|(2a_n\log(N_n/\Lambda))^{-1/2}(\alpha_n(G(x)+\Lambda h_n)-\alpha_n(G(x)))-f\|\leq \frac{\varepsilon}{8}.$$

Hence, (3.13) entails with probability 1 for any n large enough

$$(3.14) \quad \left\{ (2a_n \log(N_n/\Lambda))^{-1/2} \left(\alpha_n(G(x+h_n)) - \alpha_n(G(x)) \right) : x \in \widetilde{R} \right\} \subseteq \mathcal{S}^{5\varepsilon/8}.$$

Next we use (3.7) and (3.14) to obtain for any n sufficiently large

$$\frac{n^{1/2}h_n}{(2a_n\log(N_n/\Lambda))^{1/2}} \max_{x\in\widetilde{R}} \pm (\widetilde{f}_n(x) - \mathbb{E}\widetilde{f}_n(x))$$

$$\leq \frac{5\varepsilon}{8} \int_0^1 |d\widetilde{K}(u)| + \sup_{f\in\mathcal{S}} \mp \int_0^1 f(u)d\widetilde{K}(u).$$

We use (K2), (K3), and Schwarz's inequality to obtain

(3.16)
$$\left(\int_0^1 f(u)d\widetilde{K}(u)\right)^2 \le \left(\int_0^1 \widetilde{K}^2(u)du\right).$$

Moreover, for any $\varepsilon' > 0$ sufficiently small, (3.11) entails for any n large enough

$$(3.17) \qquad \left(\frac{\Lambda}{a(x)}\right)^{1/2} \left(1 + \frac{\log(1/\Lambda)}{\log N_n}\right)^{1/2} \le \left(1 - \frac{\varepsilon^2}{32}\right)^{-1} (1 + \varepsilon)^{1/2} \le 1 + \varepsilon'.$$

Therefore, (2.69)-(2.71) imply almost surely for any n large enough,

$$\left(\frac{nh_n}{2\log N_n}\right)^{1/2} \max_{x\in\widetilde{R}} \pm \left(\frac{\widetilde{f}_n(x) - \mathbb{E}\widetilde{f}_n(x)}{\sqrt{g(x)}}\right) \le \frac{5\varepsilon}{8}(1+\varepsilon')$$

$$\times \int_0^1 |d\widetilde{K}(u)| + (1+\varepsilon')\left(\int_0^1 \widetilde{K}^2(u)du\right)^{1/2} + \varepsilon.$$

For $\delta_0 > 0$ sufficiently small, R can be seen as a finite union of sets \widetilde{R}_i [similar to \widetilde{R}] with $1 \le i \le L < \infty$. Therefore, using the same proof as the one developed earlier, inequality (3.18) can be obtained for each \widetilde{R}_i , $1 \le i \le L$, and as a consequence, for any $x \in R$. Next, we let ε and ε' converge to 0 to obtain (3.2).

Proof of (ii). The proof of (3.3) is very similar to the one developed before and therefore omitted.

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